

Gender and the Family

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14.771

Is the household efficient? Take 2-consumption

- Lack of production efficiency is bad news for efficiency.
- On the consumption side, empirical implication of full commitment should be that, conditional on total household demand, the demand for good k should be invariant across states.
- Alternatively: the shocks should affect the consumption of each individual good only to the extent that they affect *total* consumption.
- Intuition: Husband and wife should insure each other completely. So her consumption of favorite item should not drop because she got a bad draw; of course, total consumption will change and hence basket consumed will change due to income effect, but conditioning on total expenditure we should not see an effect of the shares .

Experimental version: Robinson, 2012

- The experiment followed 142 married couples for 8 weeks in Kenya
- Every week, each individual had a 50 percent chance of receiving a 150 Kenyan shilling (KS) (US \$ 2.14) income shock, equivalent to roughly 1.5 days' income for men and 1 week's income for women
- Information about the shocks was public knowledge: both spouses were told what their partner received.
- Weekly data on consumption, income and income shocks and labor supply from each member

Predictions and Results

- shocks are, random, transitory, and idiosyncratic : Should not affect bargaining power
- Public: no moral hazard or anything....
- Shocks are spent privately [▶ Table](#)
- And saved privately [▶ Table](#)

Non experimental version: Duflo Udry

- Setting: Cote D'Ivoire
- Women and Men (tend to) grow different crop, on their different farm.
- A special crop is Yam, which is to be used by men for household public goods.
- We can compute proxies for male and female income (and yam income) by aggregating crop income across different crops.

Duflo Udry

- We first predict $y_{si2} - y_{si1}$, for s in $\{m, f, y\}$ as a function of rainfall and form predicted value of those difference

$\hat{\Delta}y_{si} = y_{si2} - y_{si1}$, and we run

$$\Delta(\log(c_i)) = \alpha + \beta \hat{\Delta}y_{fi} + \gamma \hat{\Delta}y_{mi} + \delta \hat{\Delta}y_{yi} + \epsilon_i$$

in a Pareto-efficient model, why would the coefficient β , γ and δ differ?

- What test of Pareto-efficiency does this suggest?
- Consumption of particular goods should change only to the extent that total expenditure changes.
- Two steps:

- 1 Run the same regression with total expenditures are the dependent variable

$$\Delta(\log(x_i)) = \pi_1 + \pi_2 \hat{\Delta}y_{fi} + \pi_3 \hat{\Delta}y_{mi} + \pi_4 \hat{\Delta}y_{yi} + \epsilon_i$$

- 2 calculate the ratios: $\frac{\beta}{\pi_1}$, $\frac{\gamma}{\pi_2}$, $\frac{\delta}{\pi_3}$. They should all be equal.
- 3 They are not [▶ Results](#)

Taking stock

- Consumption risk does not appear to be efficiently shared
- Especially idiosyncratic production shocks
- The standard EHM is not doing very well.
- Why not?

“Modern” family and gender economics

- “Modern” family and gender economics goes beyond testing household models and tries to understand the root of different gender preferences and the actual way in which households bargain (or not)
- No replacement model has really emerged, but lots of interesting things.

Information flow in the household

- Perhaps the premise that households member would share information efficiently is not correct.
- Rao, Ridley, Schilbach (2021): Husband and Wife do not share information that they have. Especially, husband do not pay attention to what wife may know.

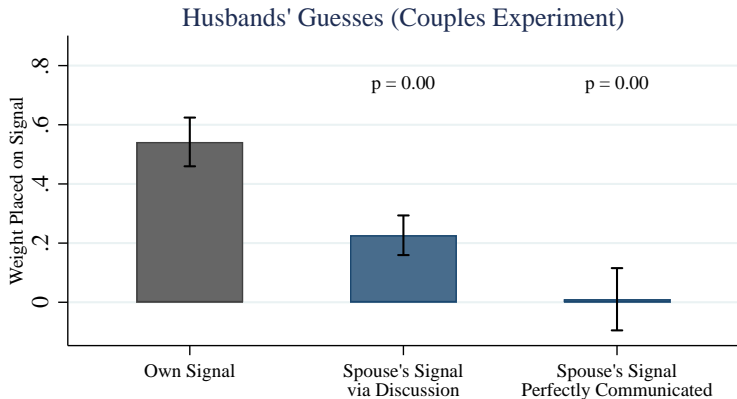
Social learning experiments with 400 couples and 500 strangers in Chennai

- 1 Do people respond similarly to info uncovered by themselves and by their spouse?
- 2 Does this vary by gender?
- 3 Is inefficient learning due to a lack of communication or incorrect weighting of info?
- 4 Are strangers working in teams similar to spouses?

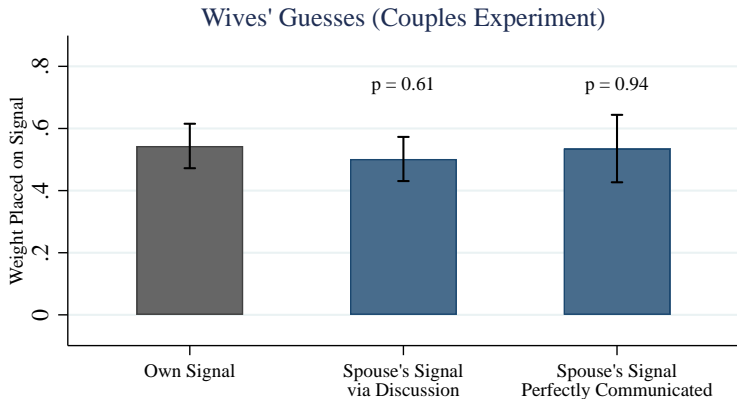
Simple, incentivized task: draw signals and guess share of red balls in an urn

- Experimental variations
 - (1) Draw all signals privately yourself
 - (2) Some signals drawn by your spouse — can learn via discussion
 - (3) Some signals drawn by your spouse — directly inform you of spouse's signals
- Key outcome of interest: weight put on signals depending on who uncovered them
- Clear prediction for information-pooling: treat own and spouse's info equally

Husbands heavily discount wife's info (even if perfectly communicated)



Wives treat own and husband's info the same



What happens with strangers?

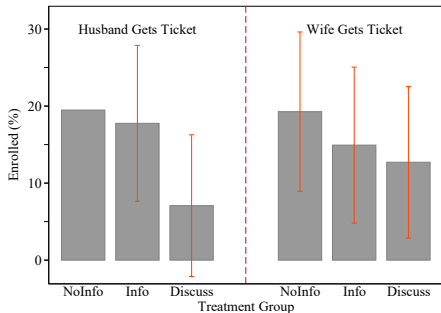
- People discount stranger information
- This time both men and women do!

Manipulation of information

Ashraf Field and Lee

- Women are proposed a voucher to jump the queue to get an injectable contraceptive.
- Two treatments: individual or couple.
- The individual one is much more likely to lead to contraceptive use than a couple intervention [▶ Ashraf et al, 2014](#)
- Not necessarily a generalizable results. Lowe and McKelway (2018): Men do not need to manipulate information in India, when information on a job opportunity is given to them in private it makes no difference when it is given to them with knowledge of their spouse. Forcing people to negotiate together reduces take up of the job opportunity [two unexpected results!]

Low and McKelway (2018)



Female Labor supply and empowerment

- In most developed country settings, female labor supply is considered as a marker for low bargaining power (leisure=private good)
- But in many developing country settings (perhaps particularly in South Asia) it seems women want to work and their husband do not want them to work (Fletcher, Pande, Moore 2019)
- “Acting wife” : in a very different context (women attending MBA at top B school) , unmarried women were willing to take costly steps to not demonstrate professional ambition in front of men (Bursztyn, Fujiwara, Pallais, 2017)
- This is consistent with limited commitment EHM: women want to work to increase their bargaining power, and men don't want that, either to protect their own bargaining power, or because they have direct disutility to see their woman work
- There could also be a social norm against female working.

Culture and Social norms

Alesina, Guiliano and Nunn

- Esther Boserup's hypothesis: in regions where the plow was dominant, males were more involved in the working of the field, and women less valued
- Hypothesis: this persisted over time.
- They use FAO data base on crop suitability to build an index of where the plow was more likely to be used.
- And correlated with today's social norms.

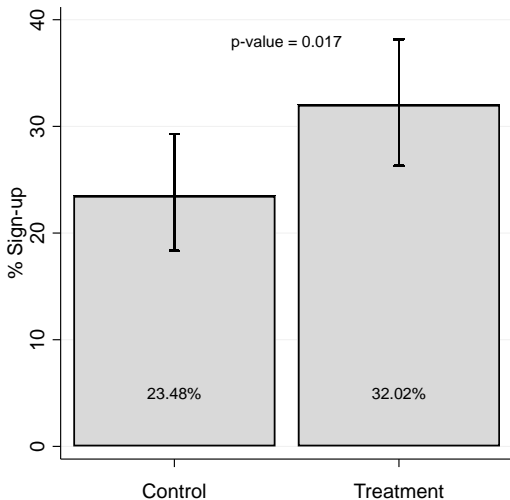
Miss-perception of social norms

Burztein, Gonzalez and Yanagizawa-Drott

- Experiment with 500 young men in Saudi Arabia
- 87% agree with the statement “In my opinion, women should be allowed to work outside of the home”
- But when asked how many other men have this opinion, three quarter under estimate the true number
- The experiment gives half of them the right number.
- Then they got the choice between \$5 Amazon certificate and opportunity to sign their wife for a platform on job.
- And follow up calls for longer term outcomes.

Miss-perception of social norms

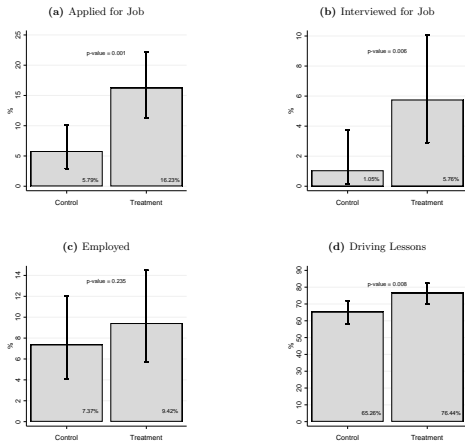
Figure 4: Job-Matching Service Sign-up
(*Main Experiment*)



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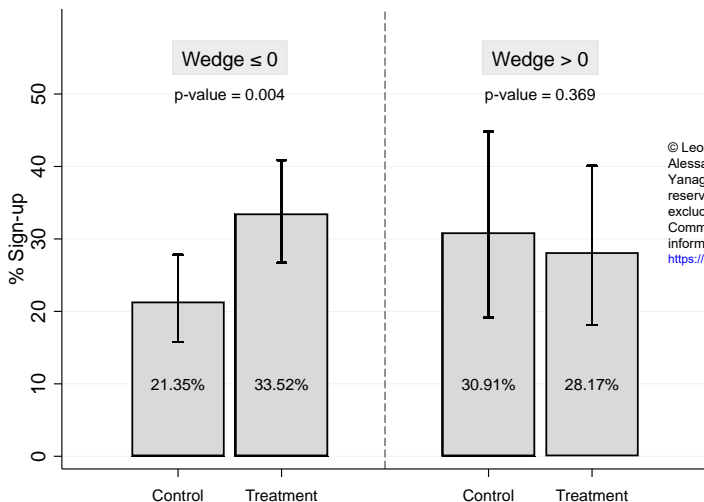
Miss-perception of social norms

Figure 5: Long-term Labor Supply Outcomes
(Follow-up)



Miss-perception of social norms

Figure 6: Job-Matching Service Sign-up–Heterogeneity by Wedge
(*Main Experiment*)



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Can Norms be changed by teaching?

Dhar, Jain, Jayachandran "Reshaping adolescents' gender attitude"

- This mis-perception suggests that perhaps norms are not a fatality
- A literature shows that relatively superficial interventions change norms such as whether females can make good leaders (Beaman et al, 2013), fertility (La Ferrara, soap opera in Brazil).
- Work in collaboration with a local NGO in North India (Breakthrough) to try to affect adolescent view of women and girls
- 45 minutes classroom discussions on various topic related to gender once every 3 weeks for 2 school years
- RCT in 314 schools in Haryana (a state in India with very bad gender culture), 14,000 students
- Find 0.25 SD improvement on self-reported gender norms at end of intervention, and some effects on behavior (especially among boys).

Teaching social norms

Table 2: Average effects of the gender attitude-change intervention

| | Gender Attitudes Index (1) | Aspirations Index (2) | Girls' Behavior Index (3) | Boys' Behavior Index (4) | Behavior Index (5) |
|-------------------|-------------------------------------|-----------------------------|------------------------------------|-----------------------------------|--------------------------|
| Treated | 0.250*** [0.019] | 0.052*** [0.019] | 0.199*** [0.031] | 0.461*** [0.031] | 0.323*** [0.022] |
| Basic controls | Yes | Yes | Yes | Yes | Yes |
| Extended controls | No | No | No | No | No |
| Observations | 13988 | 13988 | 7787 | 6201 | 13988 |

Getting women to participate in the labor force

- Given the disagreement between men and women on labor supply one could:
 - ① Change husband's opinions
 - ② Change wife's ability to advocate for themselves

McKelway, 2021 “Women’s employment in India: Intra-household and intra personal constraints

- Job market paper: experiment she conducted on her own on a shoestring...well worth reading!!
- Cross randomized two interventions with large carpet manufacturer in India who was interested in recruiting more women.
- Setting: Uttar Pradesh, poor area with backwards gender norms and very low FLP
 - Psychosocial intervention (Generalized Self Efficacy, Bandura 1977)-training over several weeks
 - Promotion of the job to the husband and in laws (6 minutes video)

GSE training affect GSE, not promo

Table 3: Effects on Women's GSE

| | (1) | (2) | (3) | (4) |
|---|------------------------------|-----------------------------|-----------------------------|------------------------------|
| | % GSE Questions Agreed With | | | |
| | at 5 Weeks | at 6 Weeks | at 5 Months | at 13 Months |
| Panel A: Unsaturated Specification | | | | |
| γ_1 : GSE Treat | 4.959 (2.013) [0.015] | 3.230 (1.796) [0.074] | 3.123 (1.681) [0.065] | 3.890 (1.964) [0.049] |
| γ_2 : Promo Treat | 1.548 (2.135) [0.469] | 0.121 (1.938) [0.950] | 0.337 (1.794) [0.851] | 0.032 (2.211) [0.988] |
| P-Value for Test that: | | | | |
| $\gamma_1 = \gamma_2$ | 0.254 | 0.228 | 0.240 | 0.197 |
| Strata FE | Yes | Yes | Yes | Yes |
| PDS Lasso X | Yes | Yes | Yes | Yes |
| Panel B: Saturated Specification | | | | |
| β_1 : GSE Treat & Promo Control | 2.662 (3.022) [0.379] | 5.607 (2.758) [0.043] | 3.616 (2.505) [0.150] | 3.286 (2.741) [0.232] |
| β_2 : GSE Control & Promo Treat | -1.130 (2.977) [0.705] | 2.344 (2.590) [0.366] | 0.625 (2.513) [0.804] | -1.182 (2.981) [0.692] |
| β_3 : GSE Treat & Promo Treat | 6.638 (2.825) [0.019] | 3.231 (2.684) [0.229] | 3.330 (2.597) [0.201] | 3.271 (2.865) [0.255] |
| P-Value for Test that: | | | | |
| $\beta_1 = \beta_2$ | 0.196 | 0.197 | 0.208 | 0.127 |
| $\beta_1 = \beta_3$ | 0.152 | 0.364 | 0.906 | 0.996 |
| $\beta_2 = \beta_3$ | 0.004 | 0.716 | 0.276 | 0.131 |
| Strata FE | Yes | Yes | Yes | Yes |
| PDS Lasso X | Yes | Yes | Yes | Yes |
| GSE Control & Promo Control Mean | 72.166 | 75.093 | 78.081 | 73.793 |
| N Women | 868 | 855 | 795 | 674 |

GSE and promo alone affect work off farm, but not combined

Table 4: Effects on Women's Employment

| | (1) Participation in Firm's Program | | (3) | (4) | (5) |
|---|--|---------------------|-------------------|---------------------------|-------------------|
| | Attended in | | at 6 Weeks | Working off Own Farm (=1) | |
| | Signed Up (=1) | First 2 Months (=1) | | at 5 Months | at 13 Months |
| Panel A: Unsaturated Specification | | | | | |
| γ_1 : GSE Treat | -0.008 (0.025) | -0.003 (0.018) | 0.001 (0.020) | 0.022 (0.024) | -0.006 (0.029) |
| | [0.739] | [0.866] | [0.963] | [0.365] | [0.847] |
| γ_2 : Promo Treat | 0.038 (0.028) | 0.016 (0.020) | -0.002 (0.022) | 0.018 (0.027) | -0.008 (0.032) |
| | [0.177] | [0.416] | [0.916] | [0.505] | [0.799] |
| P-Value for Test that: | | | | | |
| $\gamma_1 = \gamma_2$ | 0.240 | 0.496 | 0.919 | 0.921 | 0.952 |
| Strata FE | Yes | Yes | Yes | Yes | Yes |
| PDS Lasso X | Yes | Yes | Yes | Yes | Yes |
| Panel B: Saturated Specification | | | | | |
| β_1 : GSE Treat & Promo Control | 0.048 (0.035) | 0.035 (0.024) | 0.061 (0.029) | 0.087 (0.033) | 0.005 (0.043) |
| | [0.169] | [0.152] | [0.039] | [0.009] | [0.905] |
| β_2 : GSE Control & Promo Treat | 0.096 (0.038) | 0.055 (0.025) | 0.061 (0.028) | 0.076 (0.034) | -0.007 (0.041) |
| | [0.012] | [0.032] | [0.030] | [0.027] | [0.863] |
| β_3 : GSE Treat & Promo Treat | 0.029 (0.036) | 0.013 (0.025) | -0.003 (0.027) | 0.038 (0.036) | -0.012 (0.043) |
| | [0.412] | [0.605] | [0.912] | [0.301] | [0.779] |
| P-Value for Test that: | | | | | |
| $\beta_1 = \beta_2$ | 0.218 | 0.479 | 0.984 | 0.772 | 0.764 |
| $\beta_1 = \beta_3$ | 0.593 | 0.410 | 0.032 | 0.188 | 0.678 |
| $\beta_2 = \beta_3$ | 0.087 | 0.132 | 0.026 | 0.321 | 0.905 |
| Strata FE | Yes | Yes | Yes | Yes | Yes |
| PDS Lasso X | Yes | Yes | Yes | Yes | Yes |
| GSE Control & Promo Control Mean | | | | | |
| | 0.221 | 0.085 | 0.131 | 0.187 | 0.190 |
| N Women | | | | | |
| | 1022 | 1022 | 854 | 794 | 674 |

Notes: This table presents effects on women's employment. The outcome in column (1) is an indicator for signing up for the firm's program, and the outcome in column (2) is an indicator for ever attending the program in the first two months of training. The outcomes in columns (3)-(5) are indicators for having done any work for income off one's household's farm in the preceding two weeks. They come from women's six-week, five-month, and 13-month surveys. See Appendix Section G.3 for additional information on the outcomes in this table. Standard errors (in parentheses) are clustered by household and by meeting group in Panel A, and clustered by household and by meeting group \times promotion treatment in Panel B. P-values are in brackets.

Does labor supply indeed increase bargaining power?

Field, Moore, Pande, Rigol, Schaner, 2019 “on her account..”

- Experiment in Madhya Pradesh
- Government gave women access to bank account to randomly selected GP
- In one treatment they linked NREGA (workfare) payment to it
- Can therefore look at the effect of an account, and the effect of having wages linked to an account
- In the short run this increased labor supply in the program but also outside the program (including in cash payment work)
- Effects are stronger among women who had never worked for NREGA at baseline (and whose husband generally were less likely to support women working): they interpret this as increase in bargaining power

Table 3: Impact of Treatments on Women's Labour Supply

| | Aggregate Labor Supply Index | | MGNREGS Labor Supply Sub-Index | | Private Labor Supply Sub-Index | |
|--|------------------------------|---------------------|--------------------------------|-------------------|--------------------------------|---------------------|
| | Short-Run (1) | Long-Run (2) | Short-Run (3) | Long-Run (4) | Short-Run (5) | Long-Run (6) |
| <i>Panel A: Full sample</i> | | | | | | |
| θ : Direct Deposit and Training | 0.165*** (0.042) | 0.045 (0.048) | 0.186*** (0.071) | 0.021 (0.080) | 0.166*** (0.050) | 0.048 (0.062) |
| Accts Only Mean | 0.000 | -0.000 | -0.000 | 0.000 | 0.000 | -0.000 |
| N | 2504 | 2464 | 2504 | 2464 | 2504 | 2464 |
| <i>Panel B: Constrained Women</i> | | | | | | |
| θ : Direct Deposit and Training | 0.213*** (0.051) | 0.193*** (0.060) | 0.263** (0.111) | 0.069 (0.073) | 0.226*** (0.059) | 0.279*** (0.097) |
| Accts Only Mean | -0.122 | -0.186 | -0.049 | -0.102 | -0.163 | -0.275 |
| N | 922 | 903 | 922 | 903 | 922 | 903 |
| <i>Panel C: Unconstrained Women</i> | | | | | | |
| θ : Direct Deposit and Training | 0.150*** (0.052) | -0.036 (0.057) | 0.168** (0.071) | -0.008 (0.102) | 0.153** (0.071) | -0.094 (0.059) |
| Accts Only Mean | 0.061 | 0.108 | 0.033 | 0.067 | 0.080 | 0.156 |
| N | 1519 | 1501 | 1519 | 1501 | 1519 | 1501 |
| P-value: Panel B θ = Panel C θ | 0.276 | 0.001*** | 0.343 | 0.398 | 0.352 | 0.000*** |

Notes: Robust standard errors clustered at the GP level in parentheses. All regressions include strata and district fixed effects. Additional covariates are selected using double post lasso. The set of potential controls includes individual and GP-level characteristics and their square. See Online Data Appendix for the complete list of potential controls. * $p \leq 0.10$, ** $p \leq 0.05$, *** $p \leq 0.10$. The labor supply index is an average of the MGNREGS, private, and general labor sub-indices. All sub-index components are standardized with respect to the Accounts Only group. The MGNREGS

Table 4: Impact of Treatments on Empowerment

| | Aggregate Empowerment Index | | Purchase Index | | Mobility in Past Year | | Self-Reported Decision Making | |
|--|-----------------------------|---------------------|---------------------|---------------------|-----------------------|--------------------|-------------------------------|-------------------|
| | Short-Run | Long-Run | Short-Run | Long-Run | Short-Run | Long-Run | Short-Run | Long-Run |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) |
| <i>Panel A: Full sample</i> | | | | | | | | |
| θ : Direct Deposit and Training | 0.041 (0.032) | 0.032 (0.034) | 0.096* (0.053) | 0.039 (0.063) | 0.037 (0.036) | 0.053 (0.035) | -0.021 (0.053) | 0.019 (0.045) |
| Accts Only Mean | 0.000 | 0.002 | 0.000 | 0.000 | -0.000 | 0.000 | 0.000 | -0.000 |
| N | 2504 | 2453 | 2504 | 2453 | 2504 | 2464 | 2504 | 2464 |
| <i>Panel B: Constrained Women</i> | | | | | | | | |
| θ : Direct Deposit and Training | 0.100*** (0.037) | 0.144*** (0.049) | 0.239*** (0.067) | 0.238*** (0.080) | 0.023 (0.052) | 0.115** (0.056) | 0.041 (0.064) | 0.062 (0.078) |
| Accts Only Mean | -0.028 | -0.111 | -0.089 | -0.218 | 0.054 | -0.042 | -0.050 | -0.084 |
| N | 922 | 897 | 922 | 897 | 922 | 903 | 922 | 903 |
| <i>Panel C: Unconstrained Women</i> | | | | | | | | |
| θ : Direct Deposit and Training | 0.026 (0.041) | -0.022 (0.036) | 0.042 (0.065) | -0.059 (0.069) | 0.060 (0.044) | -0.001 (0.040) | -0.030 (0.071) | -0.005 (0.056) |
| Accts Only Mean | 0.010 | 0.055 | 0.037 | 0.102 | -0.031 | 0.027 | 0.025 | 0.035 |
| N | 1519 | 1496 | 1519 | 1496 | 1519 | 1501 | 1519 | 1501 |
| P-value: Panel B θ = Panel C θ | 0.145 | 0.002*** | 0.029** | 0.002*** | 0.538 | 0.061* | 0.430 | 0.487 |

Long run effect

- In the longer run norms themselves got affected. Actual norms

Table 5: Impact of Treatments on Actual Norms

| | Female Reports | | | | Male Reports | | | |
|--|------------------------|--------------------------|-------------------------------|--------------------------|------------------------|--------------------------|-------------------------------|--------------------------|
| | Actual Norms Index (1) | Personal Preferences (2) | Acceptance: Working Women (3) | Acceptance: Husbands (4) | Actual Norms Index (5) | Personal Preferences (6) | Acceptance: Working Women (7) | Acceptance: Husbands (8) |
| <i>Panel A: Full sample</i> | | | | | | | | |
| θ : Direct Deposit and Training | 0.110*** (0.040) | 0.098** (0.044) | 0.091 (0.061) | 0.087 (0.060) | -0.011 (0.043) | -0.059 (0.070) | 0.015 (0.051) | -0.024 (0.057) |
| Accts Only Mean | -0.000 | 0.000 | 0.000 | -0.000 | 0.077 | 0.180 | 0.001 | 0.049 |
| N | 2464 | 2464 | 2464 | 2464 | 2293 | 2293 | 2293 | 2293 |
| <i>Panel B: Constrained Women</i> | | | | | | | | |
| θ : Direct Deposit and Training | 0.215*** (0.051) | 0.160** (0.069) | 0.243*** (0.073) | 0.210*** (0.077) | -0.036 (0.082) | 0.012 (0.109) | -0.020 (0.083) | -0.099 (0.103) |
| Accts Only Mean | -0.095 | -0.068 | -0.099 | -0.117 | 0.066 | 0.091 | 0.045 | 0.062 |
| N | 903 | 903 | 903 | 903 | 837 | 837 | 837 | 837 |
| <i>Panel C: Unconstrained Women</i> | | | | | | | | |
| θ : Direct Deposit and Training | 0.050 (0.054) | 0.059 (0.059) | 0.007 (0.079) | 0.019 (0.073) | -0.001 (0.043) | -0.083 (0.079) | 0.040 (0.063) | -0.007 (0.057) |
| Accts Only Mean | 0.048 | 0.043 | 0.048 | 0.054 | 0.080 | 0.218 | -0.024 | 0.046 |
| N | 1501 | 1501 | 1501 | 1501 | 1403 | 1403 | 1403 | 1403 |
| P-value: Panel B θ = Panel C θ | 0.024** | 0.269 | 0.017** | 0.040** | 0.688 | 0.427 | 0.564 | 0.409 |

Notes: Robust standard errors clustered at the GP level in parentheses. All regressions include strata and district fixed effects. Additional covariates are selected using double post lasso. The set of potential controls includes individual and GP-level characteristics and their square. See Online Data Appendix for the complete list of potential controls. * $p \leq 0.10$, ** $p \leq 0.05$, *** $p \leq 0.01$. These questions were only asked in the long run survey. The actual norms index is the average of the personal preference, acceptance of working women, and acceptance of husbands sub-indices (columns 2-4). All sub-index components are standardized with respect to the Accounts Only group. The personal preferences index includes if the respondent believes that women can work, if prefers to have a daughter-in-law who wants to work for pay, and if prefers to have a son-in-law who allows daughter to work for pay. The acceptance indices are derived from a series of vignette questions featuring a housewife and working woman. The acceptance of working women sub-index includes if the respondent believes the working woman is the better wife, if believes the working woman is the better mother, and if believes the working woman is the better caretaker. The acceptance of husbands index includes if the respondent believes the working woman's husband is a better provider and if believes the working woman's husband is a better husband. See Online Data Appendix for further details on variable construction.

Long run effect

- In the longer run norms themselves got affected. Perceived norms.

Table 6: Impact of Treatments on Perceived Norms

| | Female Reports | | | Male Reports | | |
|--|-----------------------------|---|--|---------------------------|---|--|
| | Perceived Norms Index (1) | Perceived Norms: Acceptance Working Women (2) | Perceived Norms: Acceptance Husbands (3) | Perceived Norms Index (4) | Perceived Norms: Acceptance Working Women (5) | Perceived Norms: Acceptance Husbands (6) |
| | <i>Panel A: Full sample</i> | | | | | |
| θ : Direct Deposit and Training | 0.062 (0.039) | 0.079** (0.040) | 0.050 (0.046) | 0.087** (0.044) | 0.062 (0.053) | 0.113** (0.052) |
| Accts Only Mean | -0.000 | -0.000 | 0.000 | -0.236 | -0.138 | -0.334 |
| N | 2464 | 2464 | 2464 | 2292 | 2292 | 2292 |
| <i>Panel B: Constrained Women</i> | | | | | | |
| θ : Direct Deposit and Training | 0.116* (0.069) | 0.096 (0.071) | 0.152* (0.080) | 0.102 (0.078) | 0.030 (0.095) | 0.174** (0.084) |
| Accts Only Mean | -0.079 | -0.064 | -0.094 | -0.310 | -0.188 | -0.432 |
| N | 903 | 903 | 903 | 836 | 836 | 836 |
| <i>Panel C: Unconstrained Women</i> | | | | | | |
| θ : Direct Deposit and Training | 0.007 (0.043) | 0.052 (0.048) | -0.037 (0.051) | 0.115** (0.046) | 0.090 (0.063) | 0.121** (0.052) |
| Accts Only Mean | 0.047 | 0.041 | 0.053 | -0.200 | -0.117 | -0.284 |
| N | 1501 | 1501 | 1501 | 1403 | 1403 | 1403 |
| P-value: Panel B θ = Panel C θ | 0.168 | 0.606 | 0.041** | 0.882 | 0.597 | 0.520 |

Notes: Robust standard errors clustered at the GP level in parentheses. All regressions include strata and district fixed effects. Additional covariates are selected using double post lasso. The set of potential controls includes individual and GP-level characteristics and their square. See Online Data Appendix for the complete list of potential controls. * $p \leq 0.10$, ** $p \leq 0.05$, *** $p \leq 0.01$. These questions were only asked in the long run survey. The perceived norms index is the average of the perceived acceptance of making money and the perceived acceptance of making women's husbands.

Less sanguine results from McKelway, 2020

- She follow decision making in households where husband where given promotion
- In the short run, women spend more time working but just as much time on chore
- At 4 months, they think they have more decision making power, but their family does not...
- Women quickly dropped out of the job (often because it was incompatible with her other responsibilities).
- In The Fletcher et al paper, many woman who are currently not working would consider a part time job.

McKelway, impact of labor supply on family decision

Table 3: Effects on Women's Involvement in Household Decision-Making

| | (1) | (2) |
|--------------|-----------------------------|-----------------------------|
| | Woman Makes Decisions Index | |
| | Woman's Report | Family's Report |
| Promo Treat | 0.246 (0.094) [0.009] | 0.018 (0.095) [0.849] |
| Strata FE | Yes | Yes |
| PDS Lasso X | Yes | Yes |
| Control Mean | 0.000 | 0.000 |
| N Women | 390 | 379 |

Notes: The outcomes are from the four-month end-line surveys. Respondents were asked who in their households usually makes decisions about nine different things. I define indicators that take the value of one if the woman was said to make the decision alone or together with others, and zero otherwise. I aggregate the indicators into summary indices. The outcome in column (1) is the index of women's reports, and the outcome in column (2) is the index of family members' reports. Standard errors are clustered by household

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TABLE 3
 OLS FIXED-EFFECT ESTIMATES OF THE DETERMINANTS OF PLOT YIELD AND Ln(Plot Output) (\times 1,000 FCFA)
 Dependent Variable: Value of Plot Output/Hectare

| | HOUSEHOLD-YEAR-CROP EFFECTS: ALL CROPS (1) | HOUSEHOLD-YEAR EFFECTS | | HOUSEHOLD-CROP-YEAR EFFECTS | |
|----------------------------|--|------------------------|----------------------|-----------------------------|---------------------------|
| | | Millet Only (2) | White Sorghum (3) | Vegetables (4) | All Crops: CES* (5) |
| Mean of dependent variable | 89 | 31 | 41 | 134 | 167 |
| Gender: (1 = female) | -27.70 (-4.61) | -10.36 (-2.53) | -19.38 (-4.43) | -34.27 (-2.21) | -.20 (-3.56) |
| Plot size: | | | | | |
| 1st decile | 133.99 (3.50) | -28.35 (-2.67) | -17.90 (-1.92) | 237.10 (4.66) | |
| 2d decile | 69.10 (4.38) | 8.64 (.82) | 52.30 (3.16) | 63.97 (2.38) | |
| 3d decile | 63.45 (5.52) | 16.95 (1.81) | 47.68 (4.77) | 35.87 (1.52) | |
| 4th decile | 34.08 (2.88) | 9.79 (1.12) | 26.73 (3.12) | 4.21 (.18) | |
| 6th decile | -2.04 (-.29) | -.99 (-.11) | -6.38 (-1.16) | -6.65 (-.26) | |
| 7th decile | -13.44 (-1.78) | -13.01 (-1.73) | -11.31 (-1.69) | -33.54 (-.90) | |
| 8th decile | -17.23 (-2.59) | -12.97 (-1.34) | -28.58 (-4.82) | 31.04 (.73) | |
| 9th decile | -26.68 (-3.81) | -21.50 (-2.65) | -28.65 (-4.98) | | |
| 10th decile | -31.52 (-4.49) | -20.56 (-2.55) | -37.70 (-6.03) | | |
| Ln(area) | | | | | .78 (29.52) |
| Toposequence: | | | | | |
| Uppermost | -41.35 (-2.18) | 2.50 (.24) | -14.60 (-1.73) | -131.34 (-1.82) | -.46 (-2.71) |
| Top of slope | -26.35 (-1.27) | 9.53 (.96) | -11.27 (-1.47) | -121.05 (-1.85) | -.29 (-1.92) |
| Mid-slope | -24.38 (-1.19) | 5.39 (.64) | -8.62 (-1.15) | -119.68 (-1.88) | -.28 (-1.97) |
| Near bottom | -21.70 (-.90) | 4.48 (.40) | -5.36 (-.71) | -93.96 (-1.30) | -.18 (-1.27) |

LEAST-SQUARES TOBIT FIXED-EFFECT ESTIMATES OF THE DETERMINANTS OF PLOT INPUT INTENSITIES

| | HOUSEHOLD-YEAR-CROP EFFECTS | | | | | | | | | |
|------------------------------------|-----------------------------|---------|------------------------------|---------|-----------------------------|---------|------------------------------------|---------|-----------------------------------|---------|
| | Male Labor per Hectare (1) | | Female Labor per Hectare (2) | | Child Labor per Hectare (3) | | Nonhousehold Labor per Hectare (4) | | Manure (1,000 kg per Hectare) (5) | |
| Gender (1 = female) | -668.47 | (-9.60) | 70.23 | (1.53) | -195.46 | (-2.34) | -428.41 | (-1.70) | -16.33 | (-2.54) |
| Plot size: | | | | | | | | | | |
| 1st decile | 1,209.72 | (2.55) | 1,462.21 | (5.71) | 740.80 | (1.17) | 193.35 | (.43) | 24.79 | (2.42) |
| 2d decile | 417.18 | (3.25) | 1,131.01 | (5.82) | 143.12 | (1.11) | 487.39 | (1.28) | 7.99 | (.96) |
| 3d decile | 245.94 | (2.74) | 799.12 | (6.72) | 133.16 | (1.53) | 689.39 | (1.27) | 2.58 | (.48) |
| 4th decile | 96.53 | (1.71) | 407.87 | (5.02) | 72.51 | (.68) | 378.18 | (1.07) | -6.18 | (-1.12) |
| 6th decile | - .55 | (-.01) | -69.25 | (-1.36) | -72.15 | (-.98) | 57.48 | (.80) | -2.14 | (-.33) |
| 7th decile | -153.12 | (-2.97) | -306.51 | (-5.96) | -59.53 | (-.60) | 65.51 | (.64) | -11.08 | (-1.54) |
| 8th decile | -375.53 | (-6.23) | -386.78 | (-6.61) | -184.61 | (-1.61) | -43.81 | (-.30) | -11.01 | (-1.61) |
| 9th decile | -413.36 | (-6.79) | -373.57 | (-5.16) | -269.99 | (-1.83) | -255.15 | (-1.87) | -11.64 | (-1.80) |
| 10th decile | -490.11 | (-7.72) | -418.06 | (-6.08) | -219.27 | (-1.86) | -220.64 | (-1.07) | -16.41 | (-2.45) |
| Toposequence: | | | | | | | | | | |
| Uppermost | 41.62 | (.35) | -1.92 | (-.02) | -55.52 | (-.51) | 20.20 | (.12) | -9.22 | (-.62) |
| Top of slope | 29.36 | (.30) | 91.02 | (1.07) | 35.15 | (.38) | 144.02 | (.83) | .26 | (.02) |
| Mid-slope | 36.08 | (.38) | .57 | (.01) | .10 | (.00) | -15.45 | (-.11) | 1.14 | (.11) |
| Near bottom | 16.42 | (.18) | 75.94 | (.86) | -98.03 | (-1.05) | 23.27 | (.17) | 2.88 | (.27) |
| Soil Types: | | | | | | | | | | |
| 3 | 103.49 | (.60) | -31.68 | (-.23) | 235.74 | (.86) | 175.29 | (.50) | -11.80 | (-1.18) |
| 7 | -65.79 | (-.85) | -30.39 | (-.28) | 21.88 | (.44) | 66.04 | (.47) | -.07 | (-.01) |
| 11 | -28.77 | (-.09) | -52.06 | (-.34) | -778.86 | (-4.36) | 262.71 | (.70) | -.70 | (-.08) |
| 12 | 1,051.98 | (.82) | 367.34 | (1.63) | 62.36 | (.44) | 368.47 | (1.13) | 16.32 | (1.48) |
| 13 | 274.48 | (1.33) | -38.50 | (-.29) | | | -187.07 | (-.89) | | |
| 21 | 196.37 | (.95) | -43.41 | (-.49) | -42.87 | (-.35) | 37.73 | (.27) | 2.86 | (.18) |
| 31 | 83.16 | (1.59) | 68.24 | (.92) | 205.90 | (2.29) | 115.56 | (1.00) | 6.43 | (1.29) |
| 32 | 24.77 | (.50) | -10.36 | (-.15) | 173.14 | (1.07) | -51.08 | (-.44) | .73 | (.12) |
| 33 | 250.40 | (2.57) | 163.76 | (1.36) | 206.68 | (.78) | -113.92 | (-.37) | 17.28 | (1.61) |
| 35 | 179.46 | (1.50) | 303.86 | (1.90) | 248.38 | (2.60) | 195.14 | (.58) | -12.75 | (-.94) |
| 37 | 82.49 | (.70) | 50.84 | (.30) | 114.53 | (1.19) | 31.14 | (.20) | 8.34 | (1.44) |
| 45 | 78.13 | (1.34) | -8.33 | (-.10) | 79.85 | (1.02) | 41.90 | (.25) | 8.00 | (1.83) |
| 46 | -187.14 | (-1.84) | 141.73 | (.76) | 42.70 | (.09) | 223.23 | (1.27) | -15.45 | (-1.79) |
| 51 | 95.73 | (1.83) | -27.01 | (-.33) | 2.93 | (.05) | 126.70 | (1.05) | .80 | (.17) |
| Location: | | | | | | | | | | |
| Compound | 35.35 | (.78) | 37.16 | (.90) | -18.82 | (-.31) | -162.88 | (-1.38) | .99 | (.24) |
| Village | 19.69 | (.70) | 12.18 | (.45) | 42.92 | (.93) | 25.80 | (.30) | 5.86 | (1.60) |
| Mean of dependent variable when >0 | 427.39 | | 466.18 | | 85.55 | | 84.88 | | 1.70 | |
| | 506.62 | | 517.17 | | 202.88 | | 213.11 | | 7.78 | |

NOTE.—This is the least-squares implementation of Honoré's (1992) fixed-effect Tobit estimator. *t*-ratios are in parentheses.

Table 2: First stage summary statistics

| | Dependent variables | | |
|---|---------------------|-----------------|------------------|
| | Current | | |
| | Male cash crop | Yam income | Female Income |
| | (1) | (2) | (3) |
| F statistics (p value) | | | |
| All rainfall variables are significant | 1.99 (0.014) | 3.50 (0.000) | 2.53 (0.000) |
| Current year rainfall variables significant | 1.18 (0.315) | 3.38 (0.000) | 2.43 (0.005) |
| Past year rainfall variables significant | 2.79 (0.005) | 4.64 (0.000) | 2.64 (0.001) |
| Rainfall variables significantly different from: | | | |
| Male cash crop | NA | | |
| | 2.10 (0.010) | NA | |
| Yam income | | 2.38 (0.002) | |
| Female income | | | NA |

Table 4: Restricted overidentification tests

| | Dependent variable: Change in log (item consumption) | | | | | | | | | | | |
|--|--|------------------|-------------------|-------------------|-------------------|-------------------|-------------------|------------------|-------------------|-------------------|-------------------|-----------------------|
| | Total expenditure | Food consumption | Adult goods | Clothing | Prestige goods | Education | Staples | Meat | Vegetables | Processed foods | Purchased foods | Food consumed at home |
| | (1) | (2) | (3) | (4) | (5) | (6) | (7) | (8) | (9) | (10) | (11) | (12) |
| PANEL A | | | | | | | | | | | | |
| OLS coefficients: | | | | | | | | | | | | |
| Predicted change in male non-yam income | 0.126 (0.049) | 0.062 (0.054) | 0.870 (0.425) | -0.164 (0.334) | 0.683 (0.209) | -0.101 (0.128) | 0.113 (0.072) | 0.002 (0.126) | 0.345 (0.210) | 0.004 (0.139) | -0.029 (0.078) | 0.098 (0.119) |
| Predicted change in yam income | 0.207 (0.037) | 0.227 (0.041) | -0.473 (0.320) | 0.296 (0.252) | -0.272 (0.158) | 0.320 (0.108) | 0.345 (0.054) | 0.135 (0.096) | 0.023 (0.159) | 0.122 (0.105) | 0.087 (0.059) | 0.444 (0.090) |
| Predicted change in female income | 0.309 (0.056) | 0.235 (0.061) | 1.537 (0.490) | 0.535 (0.382) | 0.993 (0.239) | -0.098 (0.159) | 0.193 (0.082) | 0.492 (0.144) | 0.995 (0.239) | 0.474 (0.159) | 0.412 (0.089) | 0.313 (0.136) |
| F tests (p value) : | | 0.934 | 5.064 | 0.514 | 7.595 | 2.260 | 5.870 | 1.824 | 3.277 | 1.397 | 4.777 | 1.912 |
| Overidentification | | (0.393) | (0.007) | (0.598) | (0.001) | (0.106) | (0.003) | (0.162) | (0.038) | (0.248) | (0.009) | (0.148) |
| Restriction test | | | | | | | | | | | | |
| PANEL B: LAGGED RAINFALL | | | | | | | | | | | | |
| OLS coefficients: | | | | | | | | | | | | |
| Predicted change in lagged male non-yam income | 0.073 (0.020) | 0.039 (0.022) | 0.350 (0.169) | 0.044 (0.133) | 0.047 (0.082) | 0.091 (0.056) | 0.038 (0.029) | 0.150 (0.050) | 0.039 (0.083) | 0.115 (0.055) | 0.155 (0.031) | -0.007 (0.047) |
| Predicted change in lagged yam income | -0.003 (0.009) | 0.004 (0.009) | 0.008 (0.073) | -0.125 (0.059) | -0.076 (0.036) | -0.031 (0.029) | -0.021 (0.013) | 0.015 (0.022) | 0.011 (0.036) | 0.027 (0.024) | 0.024 (0.013) | -0.018 (0.021) |
| Predicted change in lagged female income | -0.001 (0.026) | 0.018 (0.028) | -0.024 (0.220) | -0.251 (0.173) | -0.289 (0.107) | 0.093 (0.079) | 0.044 (0.038) | 0.023 (0.064) | -0.054 (0.107) | -0.010 (0.071) | 0.062 (0.040) | -0.035 (0.061) |
| F tests (p value) : | | 0.105 | 0.128 | 0.254 | 0.043 | 0.016 | 0.049 | 0.052 | 0.024 | 0.058 | 0.054 | 0.057 |
| Overidentification | | (0.900) | (0.880) | (0.776) | (0.958) | (0.984) | (0.952) | (0.949) | (0.976) | (0.943) | (0.948) | (0.945) |

TABLE 4
REGRESSIONS: TREATMENT EFFECTS
Dependent Variable = Dummy for Parent Prefers R\$120 CGT to R\$125 CT

| | OLS (1) | OLS (2) | OLS (3) | Probit* (4) |
|-------------------------------------|---------------------|---------------------|---------------------|---------------------|
| Text message treatment dummy | -.6136 (.087)*** | -.5819 (.091)*** | -.4744 (.097)*** | -.4634 (.072)*** |
| Don't tell treatment dummy | -.6433 (.070)*** | -.6119 (.080)*** | -.5208 (.060)*** | -.5409 (.123)*** |
| Individual and household covariates | No | Yes | Yes | Yes |
| Surveyor and school dummies | No | No | Yes | Yes |
| Observations | 156 | 156 | 156 | 151 |

NOTE.—Mean of dependent variable in the baseline group = .82. The sample was restricted to households that answered the entire survey, and thus two observations were lost. Controls in cols. 2–4 include log of household income, employed parent dummy,

TABLE 3
HOW MUCH DO PARENTS KNOW?

| Respondent | Parent | Child | Difference |
|---|--------|-------|------------|
| Did the child miss any day of school this year? (% answering "yes") | 75.60 | 85.58 | -9.98*** |
| How many days did the child miss this year? | 4.8 | 5.16 | -.36 |
| Did the child miss any day of school in the last 2 months? (% answering "yes") | 50.96 | 56.04 | -5.08 |
| How many days did the child miss in the last 2 months? | 1.36 | 1.97 | -.60* |
| Did the child miss any day this year because the child was sick? (% answering "yes") | 43.81 | 32.70 | 11.12*** |
| Did the child miss any day because the child did not want to go? (% answering "yes") | 9.05 | 15.87 | -6.82*** |

NOTE.—*T*-tests of equality in means from paired observations (parent and child).

* Significant at the 10 percent level.

*** Significant at the 1 percent level.

TABLE 3—EXPERIMENTAL SHOCKS AND EXPENDITURES

| | Expenditures | | | | | | |
|--|-------------------|---------------------|-----------------------|-------------------|-------------------|------------------------|-------------------|
| | Total (1) | Private (2) | Shared food (3) | Medical (4) | Children (5) | Other shared (6) | Transport (7) |
| <i>Panel A. Men</i> | | | | | | | |
| Shillings received in experimental shock by respondent | 0.190 (0.194) | 0.169 (0.064)*** | -0.025 (0.089) | 0.048 (0.041) | -0.012 (0.032) | -0.096 (0.102) | 0.102 (0.068) |
| Shillings received in experimental shock by spouse | -0.163 (0.192) | -0.027 (0.069) | -0.016 (0.087) | 0.057 (0.045) | -0.019 (0.030) | -0.086 (0.111) | -0.069 (0.060) |
| Observations | 898 | 898 | 898 | 898 | 898 | 898 | 898 |
| Number of households | 142 | 142 | 142 | 142 | 142 | 142 | 142 |
| <i>p</i> -value for <i>F</i> -test of equality | 0.21 | 0.05** | 0.93 | 0.84 | 0.88 | 0.95 | 0.09* |
| Mean of dependent variable (Ksh) ^a | 889.32 | 135.66 | 413.77 | 56.95 | 24.09 | 144.77 | 114.55 |
| SD of dependent variable (Ksh) | 557.30 | 122.24 | 298.74 | 143.25 | 84.40 | 250.88 | 106.76 |
| Proportion of weeks dependent variable = 0 | 0.00 | 0.12 | 0.03 | 0.52 | 0.86 | 0.12 | 0.18 |
| <i>Panel B. Women</i> | | | | | | | |
| Shillings received in experimental shock by respondent | 0.180 (0.148) | -0.020 (0.042) | 0.056 (0.067) | 0.079 (0.041)* | 0.032 (0.026) | 0.041 (0.059) | -0.007 (0.047) |
| Shillings received in experimental shock by spouse | -0.058 (0.123) | -0.026 (0.039) | -0.051 (0.064) | 0.015 (0.034) | -0.025 (0.024) | 0.050 (0.041) | -0.021 (0.039) |
| Observations | 898 | 898 | 898 | 898 | 898 | 898 | 898 |
| Number of households | 142 | 142 | 142 | 142 | 142 | 142 | 142 |
| <i>p</i> -value for <i>F</i> -test of equality | 0.14 | 0.91 | 0.23 | 0.07* | 0.1* | 0.88 | 0.77 |
| Mean of dependent variable (Ksh) | 428.51 | 47.28 | 227.98 | 28.43 | 18.25 | 68.51 | 38.07 |
| SD of dependent variable (Ksh) | 482.65 | 123.77 | 262.65 | 94.87 | 65.80 | 119.21 | 101.60 |
| Proportion of weeks dependent variable = 0 | 0.03 | 0.61 | 0.08 | 0.64 | 0.84 | 0.28 | 0.72 |

| | Net transfers to: | | Labor supply | | Savings ^b | |
|--|----------------------|-----------------------------|-------------------|------------------------|----------------------------------|-------------------------|
| | Spouse (1) | Outside household (2) | Hours (3) | Labor income (4) | Bank/ ROSCA savings (5) | Total savings (6) |
| <i>Panel A. Men</i> | | | | | | |
| Shillings received in experimental shock by respondent | 0.077 (0.065) | 0.090 (0.202) | 0.018 (0.017) | 0.139 (0.366) | 0.020 (0.159) | 0.782 (0.393)** |
| Shillings received in experimental shock by spouse | -0.163 (0.060)*** | -0.133 (0.157) | -0.036 (0.035) | -0.145 (0.312) | -0.244 (0.154) | 0.314 (0.319) |
| Observations | 898 | 898 | 898 | 898 | 898 | 898 |
| Number of households | 142 | 142 | 142 | 142 | 142 | 142 |
| <i>p</i> -value for <i>F</i> -test of equality | 0.01*** | 0.31 | 0.27 | 0.48 | 0.35 | 0.31 |
| Mean of dependent variable (Ksh) ^a | 76.78 | 2.81 | 52.18 | 698.56 | 127.20 | -270.34 |
| SD of dep. var. (Ksh) | 159.89 | 436.18 | 24.14 | 852.24 | 222.35 | 885.11 |
| Proportion of weeks dependent variable = 0 | 0.62 | 0.21 | 0.08 | 0.07 | 0.50 | 0.00 |
| <i>Panel B. Women</i> | | | | | | |
| Shillings received in experimental shock by respondent | 0.163 (0.060)*** | 0.050 (0.190) | -0.031 (0.020) | -0.020 (0.185) | 0.082 (0.088) | 0.586 (0.239)** |
| Shillings received in experimental shock by spouse | -0.077 (0.065) | -0.010 (0.160) | 0.009 (0.011) | 0.031 (0.195) | -0.154 (0.099) | 0.175 (0.234) |

TABLE 3—EFFECT OF PRIVATE INFORMATION TREATMENT ON HOUSEHOLDS IN WHICH BOTH HUSBAND AND WIFE DO NOT WANT A CHILD IN NEXT TWO YEARS

| | All women | | Responders | | Nonresponder | |
|---|-------------------------|----------------------------|-------------------------|----------------------------|-------------------------|----------------------------|
| | Voucher redeemed (1) | Received injectable (2) | Voucher redeemed (3) | Received injectable (4) | Voucher redeemed (5) | Received injectable (6) |
| <i>Panel A. Without controls</i> | | | | | | |
| Assigned to Couple treatment | -0.103** (0.049) | -0.065 (0.040) | -0.259*** (0.095) | -0.213*** (0.077) | -0.041 (0.059) | -0.014 (0.048) |
| <i>Panel B. With controls</i> | | | | | | |
| Assigned to Couple treatment | -0.097* (0.051) | -0.061 (0.041) | -0.274** (0.120) | -0.253*** (0.094) | -0.051 (0.063) | -0.020 (0.049) |
| Observations | 419 | 419 | 106 | 106 | 290 | 290 |
| Mean of outcome variable among individual treatment | 0.531 | 0.244 | 0.650 | 0.300 | 0.483 | 0.214 |

TABLE 9—EFFECTS OF TREATMENT ON SIBLINGS USING MONITORED AND ADMINISTRATIVE PARTICIPATION,
HOUSEHOLDS WITH TWO REGISTERED CHILDREN

| | Untreated children | | Female control | | Male control | |
|------------------------------------|---|----------------------|---|--------------------|--|-------------------|
| | Attendance (1) | Enrollment (2) | Attendance (3) | Enrollment (4) | Attendance (5) | Enrollment (6) |
| Sibling is treated? (yes or no) | -0.030* (0.015) | -0.073*** (0.027) | -0.053** (0.021) | -0.104* (0.054) | -0.029 (0.031) | -0.054 (0.040) |
| Child is treated? (yes or no) | | | | | | |
| Demographic controls | √ | √ | √ | √ | √ | √ |
| School fixed effects | √ | √ | √ | √ | √ | √ |
| Observations | 690 | 637 | 352 | 323 | 338 | 314 |
| R ² | 0.28 | 0.16 | 0.33 | 0.23 | 0.38 | 0.26 |
| Sample description | Untreated children in households with two registered children | | Untreated girls in households with two treated children | | Untreated boys in households with two treated children | |

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